A Comparison of APACHE II and SAPS II Scoring Systems in Predicting Hospital Mortality in Thai Adult Intensive Care Units

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Objective: To assess the performance of Acute Physiology and Chronic Health Evaluation II (APACHE II) and Simplified Acute Physiology Score II (SAPS II) in Thai critically ill patients.

Material and Method: Prospective observational cohort study conducted between July 1, 2004 and October 31, 2005 in the Intensive Care Unit (ICU) of Songklanagarind Hospital, an 800-beds tertiary referral university teaching hospital.

Results: One thousand three hundred sixteen patients were enrolled. There were 310 deaths (23.6%) at hospital discharge. APACHE II and SAPS II predicted hospital mortality 30.5 + 28.2 and 30.5 + 29.8 respectively. Both models showed excellent discrimination. The discrimination of APACHE II was better than SAPS II (0.911 and 0.888, p < 0.001). However, both systems presented a poor calibration. The Hosmer-Lemeshow goodness-of-fit H and C statistics were 66.59 and 66.65 of APACHE II (p < 0.001) and 54.01 and 71.44 of SAPS II (p < 0.001).

Conclusion: APACHE II provided better discrimination than SAPS II, but both models showed poor calibration in over predicting mortality in our ICU patients. Customized or new severity scoring systems should be developed for critically ill patients in Thailand.

Keywords: Severity of illness, Intensive care, Mortality prediction, Acute Physiology and Chronic Health Evaluation (APACHE II), Simplified Acute Physiology Score (SAPS II)

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Severity scoring systems are designed to provide an estimate of the probability of hospital mortality for critically ill patients⁽¹⁾. Measuring the severity of illness enables hospital administrators to describe ICU populations and can be helpful in clinical decision making and guiding resource allocation^{(2)}. Severityadjusted outcomes should not be used only to measure individual ICU performance, but can also be used to compare between $ICU^{(1-3)}$. Severity scoring systems have been around for many years now. The first two

introduced were the Acute Physiology and Chronic Health Evaluation score (APACHE) and the Simplified Acute Physiology Score (SAPS). Modified versions of both models have been created since then, notably APACHE $II^{(4)}$ and $III^{(5)}$, and SAPS $II^{(6)}$ and $III^{(7)}$. The most popular current scales are APACHE II and SAPS II due to their reliability and relative ease of use (1) .

Ability of any such systems to accurately predict mortality rate is determined by its discrimination and calibration. A model has good discrimination when it can distinguish accurately between patients who will die or survive. Calibration evaluates the degree of correspondence between the estimated overall probability of death produced by a model and the actual mortality rate. When the observed number of

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deaths is close to the number expected from the model, the model is considered to be well calibrated (1) . Good discrimination of both APACHE II and SAPS II has been found in several studies conducted in Western countries(8-12); however, most of these studies reported poor calibration⁽⁹⁻¹²⁾. It is not clear to what extent these findings should be extrapolated to ICU patients in different ICUs or different countries⁽¹⁾. In Asian countries, there have been four published studies in which the data showed good discrimination of APACHE $II⁽¹³⁻¹⁶⁾$, three in which SAPS II exhibited good discrimination $(13,14,16)$ and two which showed poor discrimination in APACHE II^(17,18). All studies reported a lack of fit for both systems $(13,16)$; however, only a single study reported calibration of these scores in Thailand⁽¹⁴⁾. Before applying severity scoring systems in a specific country or different type of ICU, their prognostic performance must be validated.

In the present study, the authors evaluated and compared the ability and validity of APACHE II and SAPS II scoring systems in predicting hospital mortality at a mixed medical-surgical ICU of tertiary referral university hospital in southern Thailand.

Material and Method

Patients and Data collection

The present study was performed at Songklanagarind Hospital, an 800-beds tertiary referral university teaching hospital at Prince of Songkla University, Hat Yai, Songkhla. In the hospital, there are two units in the adult ICU: a ten-bed surgical ICU and a nine-bed mixed medical and coronary care unit. The surgical ICU serves all post operative patients. The medical ICU serves all critically ill patients from the Department of Internal Medicine.

Data collection lasted from July 1, 2004 until October 31, 2005. All data were collected concurrently for consecutive ICU admissions. Patients who were excluded from the present study included those who were younger than 15 years of age, suffered burn injuries, had not received attempted cardiac resuscitation, were withdrawn from treatment, died within four hours of admission to ICU or who stayed in the ICU less than 24 hours. If patients had been admitted more than once to the ICU during the study period, only the first admission was included. Approval for the project was obtained from the faculty Ethics Committee.

The following data were collected: basic demographic characteristics including sex, age, and principal diagnostic category leading to ICU admission according to Knaus et al⁽⁴⁾. Type of patient status was defined classically as medicine, scheduled surgery, or unscheduled surgery⁽⁶⁾. The worst physiological values in the 24 hours following ICU admission for APACHE II and SAPS II variable were abstracted from clinical and laboratory records, using the variable definitions reported in the original literature $(4,6)$. The standard mortality ratio (SMR) was calculated by dividing observed hospital mortality by the predicted hospital mortality. The 95% confidence interval (CI) for the SMRs were calculated by regarding the observed mortality as a Poisson variable, then dividing its 95%CI by the predicted mortality^{(19)}. The presence of organ dysfunction during the first 24 hours in the ICU was assessed by the Sequential organ dysfunction scores (SOFA), as described by Vincent et $al^{(20)}$. In sedated patients, the Glasgow Coma Score (GCS) was determined either from medical records before sedation or through interviewing the physician who ordered the sedation. However, if a variable could not be measured the GCS was assumed normal. The predicted hospital mortality was calculated using the original formulas of each severity scoring systems^(4,6). ICU, hospital length of stay (LOS) and lead time (the interval from hospital admission to ICU admission) were calculated. Patients were followed up until ICU and hospital discharge in order to registrar their survival status. All data were collected by one well-trained research assistant and then all patients' records were rechecked for accuracy by the authors.

Statistical analysis

Stata 7 software (Stata Corporation, College Station Tx.) were used for statistical analysis. Data are presented as mean + SD, when indicated. Student's t test and Wilcoxon's rank sum test were used to compare normally distributed continuous variables and non-parametric data, respectively. Chi-square statistic was used to test for the statistical significance of categorical variables. A p*-*value less than 0.05 was considered statistically significant.

The ability and accuracy of the models for hospital mortality prediction were determined by examining their discrimination and calibrations. Discrimination was tested by the area under the receiver operating characteristics (ROC) curve, computed by a modification of the Wilcoxon statistic, as described by Hanley and McNeil⁽²¹⁾ and 2×2 classification table at decision criterion of 50%. The Hosmer-Lemeshow goodness-of-fit H and C statistics were used to evaluate calibration⁽²²⁾. Patients were rank-ordered in ten groups according to their probability of death to cal-

culate the H statistic. To calculate the C-statistic, the study population was divided into deciles of predicted risk. The predicted and actual number of survivors and non-survivors were compared statistically using formal goodness-of-fit calculations to determine whether the discrepancies between predicted and actual values were statistically significant $(p < 0.05)$.

Results

During the study period, data were collected on 1,316 patients. Overall 236patients (17.9%) died in the ICU and 310 patients (23.6%) died in the hospital. The SMR using both models was $0.77 (95\% CI = 0.69 - 0.86)$.

The patients' demographic, type of admission and general diagnostic categories of the presented patients are shown in Table 1. The severity of patients' illness as assessed by APACHE II and SAPS II are shown in Table 2.

The area under the ROC curve was 0.911 (95%CI = 0.891-0.930) for APACHE II and 0.888 (95%CI $= 0.867 - 0.909$ for SAPS II. When the APACHE II and SAPS II curves are compared, the area under the ROC curve of APACHE II showed to be statistically and

significantly, higher than SAPS II system $(p < 0.001)$ (Fig. 1).

The results of the 2×2 classification table for APACHE II and SAPS II are shown in Table 3. The false-positive rate was lower for APACHE II than SAPS II (8.45% *vs.* 10.83%).

The Hosmer-Lemeshow goodness-of-fit test H statistic revealed a poor performance for both systems (Table 4). Similar results were obtained, using the Hosmer-Lemeshow goodness-of-fit test C statistic (Table 5). These findings indicated a significant a lack of fit for both models.

Predicted risk of hospital mortality as described by the two models showed a highly significant correlation ($R = 0.907$). It is clear that, although highly correlated, the predicted mortality from the two severity scoring systems are widely dispersed (Fig. 2), mostly in the mid-range of predicted risk of death. The predictions of hospital deaths seem cloudier related in the extremes of risk.

Regarding the bias resulting from case mix, a sub sample of cases was analyzed excluding patients with coronary care and postoperative cardiac surgery

| | Total $(n = 1316)$ n (%) | Survivors $(n = 1006)$ n (%) | Non-survivors $(n = 310)$ n (%) | p-value |
|--------------------------|----------------------------------|---|---|---------|
| Age (years) | $55.6 + 18.2$ | $55.7 + 17.9$ | $55.3 + 19.1$ | 0.797 |
| Male 754 (57.3) | 570 (56.7) | 184 (59.3) | 0.402 | |
| Type of admission | | | | |
| Medicine | 732 (55.6) | 496 (49.3) | 236(76.1) | < 0.001 |
| Surgery, scheduled | 357(27.2) | 337 (33.5) | 20(6.5) | < 0.001 |
| Surgery, unscheduled | 227(17.2) | 173 (17.2) | 54 (17.4) | 0.928 |
| Categories of diseases | | | | |
| Non-operative | | | | |
| Respiratory disease | 86(6.5) | 76 (7.6) | 10(3.2) | 0.007 |
| Cardiovascular disease | 350 (26.7) | 262(26) | 88 (28.4) | 0.414 |
| Coronary artery disease | 186 (14.2) | 172(17.1) | 14(4.5) | < 0.001 |
| Others | 164(12.5) | 90(8.9) | 74 (23.8) | < 0.001 |
| Sepsis | 220 (16.7) | 103(10.2) | 117(37.7) | < 0.001 |
| Neurological disease | 16(1.2) | 13(1.3) | 3(1) | 0.649 |
| Gastrointestinal disease | 30(2.3) | 13(1.3) | 17(5.5) | < 0.001 |
| Other | 30(2.2) | 29(2.9) | 1(0.3) | 0.008 |
| Post-operative disease | | | | |
| Post-CABG* | 55 (4.2) | 48(4.8) | 7(2.2) | 0.053 |
| Post cardiac surgery | 156(11.9) | 148 (14.7) | 8(2.6) | < 0.001 |
| Brain and spinal cord | 124(9.4) | 110(10.9) | 14(4.5) | < 0.001 |
| Gastrointestinal | 94(7.1) | 78 (7.8) | 16(5.2) | 0.121 |
| Other | 155(11.8) | 126(12.5) | 29(9.4) | 0.001 |

Table 1. Demographic and clinical characteristics of patients in study

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| | Total $(n = 1316)$ | Survivors $(n = 1006)$ | Non-survivors $(n = 310)$ | p-value |
|-----------------------------------|-----------------------|----------------------------------|------------------------------|---------|
| Age (years) | $55.6 + 18.2$ | $55.7 + 17.9$ | $55.3 + 19.1$ | 0.797 |
| APS [#] | $14.7 + 9.1$ | $11.5 + 6$ | $25.1 + 9.4$ | < 0.001 |
| APACHE II | $18.3 + 9.6$ | $14.8 + 6.6$ | $29.5 + 9.3$ | < 0.001 |
| SAPS II | $40.1 + 20.4$ | $33.0 + 14.2$ | $63.1 + 20.5$ | < 0.001 |
| SOFA | $5.8 + 4.2$ | $4.5 + 3.2$ | $10.1 + 4.1$ | < 0.001 |
| APACHE II prediction of death (%) | $30.5 + 28.2$ | $19.6 + 18$ | $66.0 + 25.8$ | < 0.001 |
| SAPS II prediction of death (%) | $30.5 + 29.8$ | $19.9 + 20.6$ | $65.0 + 28.8$ | < 0.001 |
| ICU LOS (day)* | $2(1-5)$ | $2(1-4)$ | $3(1-6)$ | 0.023 |
| Hospital LOS (day)* | $15(8-29)$ | $17(10-31)$ | $8(2-22)$ | < 0.001 |
| Lead time $(\text{day})^*$ | $1(0-5)$ | $1(0-5)$ | $1(0-5)$ | 0.275 |

Table 2. APACHE II and SAPS II scores, predicted risk of hospital mortality and LOS of patients in study

APS; acute physiology scores of APACHE II

* Median and interquartile range

as in the original SAPS II and APACHE II models. Discrimination and calibrations of both systems after excluding these patients are shown in Table 6. The area under the ROC curve of APACHE II was 0.913 after excluding post coronary artery bypass graft (CABG) patients (as in the original APACHE II model). However, the Hosmer-Lemeshow goodness-of-fit test H and C statistics also showed poor calibration. The SAPS II scoring system showed good discrimination

after calculating as in the original SAPS II, but the poor goodness-of-fit persisted.

Discussion

APACHE II and SAPS II are two of the most popular models in common use throughout the world, although later version of each, APACHE III and SAPS III, have also been created. However, there are limitations in the application of these systems arising from

Fig. 1 Receiver operating characteristic curves (ROC) for APACHE II and SAPS II systems

Table 3. Classification table of APACHE II and SAPS II

| Decision criterion 50% | APACHE II | | SAPS II | | |
|--------------------------------|---------------------|-----|---|-----|--|
| | | | Predicted to live (n) Predicted to die (n) Predicted to live (n) Predicted to die (n) | | |
| Observed survivors | 921 | 85 | 897 | 109 | |
| Observed non-survivors | 81 | 229 | 91 | 219 | |
| Sensitivity | 73.87 (68.61-78.67) | | 70.65 (65.23-75.66) | | |
| Specificity | 91.55 (89.66-93.20) | | 89.17 (87.08-91.02) | | |
| Positive predictive value | 72.93 (67.65-77.78) | | 66.77 (61.38-71.85) | | |
| Negative predictive value | 91.92 (90.05-93.53) | | 90.78 (88.81-92.52) | | |
| Overall correct classification | 87.38 (85.47-89.13) | | 84.80 (82.75-86.70) | | |

* In parenthesis are 95%CI

Table 4. The Hosmer-Lemeshow goodness-of-fit test H statistic for APACHE II and SAPS II

| APACHE II predicted $\mathbf n$ | Non-survivors | | Survivors | | |
|--|---------------|----------|------------------|----------------|----------|
| death rate | | Observed | Expected | Observed | Expected |
| $0.00 - 0.10$ | 413 | 9 | 23.80 | 404 | 389.20 |
| $0.11 - 0.20$ | 268 | 16 | 38.32 | 252 | 229.68 |
| $0.21 - 0.30$ | 151 | 17 | 36.92 | 134 | 114.08 |
| $0.31 - 0.40$ | 92 | 14 | 32.33 | 78 | 59.67 |
| $0.41 - 0.50$ | 78 | 25 | 34.81 | 53 | 43.19 |
| $0.51 - 0.60$ | 66 | 30 | 36.62 | 36 | 29.38 |
| $0.61 - 0.70$ | 60 | 35 | 39.37 | 25 | 20.63 |
| $0.71 - 0.80$ | 60 | 44 | 45 | 16 | 15.00 |
| $0.81 - 0.90$ | 61 | 55 | 51.94 | 6 | 9.06 |
| $0.91 - 1.00$ | 67 | 65 | 63.07 | \overline{c} | 3.93 |
| Total | 1316 | 310 | 402.19 | 1006 | 913.81 |

Chi-square H = 66.59 *df* = 10 p < 0.001

Chi-square $H = 54.01$ *df* = 10 p < 0.001

| APACHE II | | Non-survivors | | Survivors | | |
|-------------------------|------|----------------|----------|------------------|----------|--|
| predicted death rate | n | Observed | Expected | Observed | Expected | |
| $0.00 - 0.05$ | 131 | | 3.1 | 89 | 86.9 | |
| $0.06 - 0.08$ | 132 | $\overline{4}$ | 3.4 | 80 | 80.6 | |
| $0.09 - 0.11$ | 131 | 4 | 4.5 | 89 | 88.5 | |
| $0.12 - 0.16$ | 132 | 5 | 4.2 | 66 | 66.8 | |
| $0.17 - 0.22$ | 132 | 7 | 6.7 | 80 | 80.3 | |
| $0.23 - 0.38$ | 131 | 12 | 9.6 | 71 | 73.4 | |
| $0.39 - 0.48$ | 132 | 20 | 18.1 | 64 | 65.9 | |
| $0.49 - 0.68$ | 131 | 36 | 35.5 | 48 | 48.5 | |
| $0.69 - 0.89$ | 132 | 55 | 60.3 | 33 | 27.7 | |
| $0.90 - 1.00$ | 132 | 71 | 69.6 | 9 | 10.4 | |
| Total | 1316 | 310 | 402.19 | 1006 | 913.81 | |

Table 5. The Hosmer-Lemeshow goodness-of-fit test C statistic for APACHE II and SAP II

Chi-square C = 66.65 $df = 10$ p < 0.001

Chi-square C = 71.44 $df=10$ p < 0.001

the need to validate them first using independent sample data from different ICUs in different countries such as in Thailand.

In the present study, the authors evaluated the ability and validity of APACHE II and SAPS II systems to accurately predict hospital mortality in a Thai adult mixed-case ICU. Both models showed excellent discrimination, although the authors found that discrimination was better for APACHE II than for SAPS II; the same is true for the percentage of overall correct classification. Good discrimination of both models has been reported in previous studies $(8-13,15)$. The area under the ROC of both systems in the present study

was higher than that in the other reports. Previously reported area under the ROC curve of APACHE II and SAPS II included 0.839 and 0.870 in Greece⁽¹⁰⁾. 0.787 and 0.817 in Portugal⁽¹¹⁾, 0.83 and 0.79 in Saudi Arabia⁽¹³⁾, 0.819 and 0.840 in Tunisia⁽¹²⁾ and 0.88 and 0.87 in Hong Kong^{(16)}, respectively and 0.88 in the original SAPS $II^{(6)}$. The ability of the two models to correctly predict group prognosis also was assessed by means of 2×2 decision table. Other reports using this method showed lower correct classification, for instance 85.5% in the original APACHE $II^{(4)}$, and then 83.06% and 82.75% in Greece⁽¹⁰⁾, 78.00% and 76.37% in Portugal⁽¹¹⁾, 77% and 77% in Saudi Arabia⁽¹³⁾ with

| | | APACHE II | | SAPS II | | | |
|--|------|------------------|--------------------|----------------------|---------|--------------------|-----------|
| | n | AUC^* | H chi ₂ | C chi ₂ | AUC^* | H chi ₂ | C chi 2 |
| All patients Exclude | 1316 | 0.911 | $66.59**$ | $66.65**$ | 0.888 | $54.01**$ | $71.44**$ |
| $-CABGs$ | 1261 | 0.913 | $66.69**$ | $67.18**$ | 0.891 | 54.83** | $72.35**$ |
| - $CABG + CAD^{s+}$ | 1078 | 0.915 | $51.12**$ | $50.86**$ | 0.893 | $38.23**$ | 55.78** |
| - $CABG + CAD^{s+}$ postoperative cardiac surgery | 923 | 0.904 | $44.13**$ | $44.5**$ | 0.880 | $30.77**$ | $47.33**$ |

Table 6. Discrimination and calibration of APACHE II and SAPS II after excluding coronary and postoperative cardiac surgery cases

* AUC; Area Under the Receiver Operating Curve

\$ CABG; Coronary Artery Bypass Graft

+ CAD; Coronary Artery Disease

the APACHE II and SAPS II models, respectively.

In Thailand, previous reports showed the area under the ROC curve of APACHE II include 0.723⁽¹⁸⁾, $0.788^{(17)}$ and $0.838^{(14)}$. Lertsithichi et al⁽¹⁴⁾ found the area under the ROC curve of SAPS II 0.818 in surgical patients. Limitations of previous reports were relatively small case mix populations and only a single study

reported calibration of these scores. Thus, both the APACHE II and SAPS II scoring systems are able to predict group outcome in Thai ICU patients, providing a prognostic estimate that may be helpful in assisting clinical decision making.

However, both models in the present study failed to predict hospital mortality accurately- that is,

Fig. 2 Correlation of APACHE II and SAPS II predicted hospital mortalities

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overall calibration was poor. This lack of overall goodness-of-fit is in similar to the finding of other previous studies(8,10-13,16).

Potential reasons for poor calibration of both scoring systems in the presented population might include the following: 1) differences in data collection and definition; 2) differences in case mix from other studies; 3) lead time bias. The reliability of the data collected is important because poor data can influence the predictions of mortality. Holt et $al^{(23)}$ showed that the main causes of data error in scoring APACHE II are inconsistent choice between highest and lowest value of acute physiologic score and GCS. The variability of GCS determination in sedated patients may affect the predicted death in both models. In the present study, the authors used the pre-sedation GCS in sedated patients as in previous studies (13) , an approach which has been shown to be associated with better performance of APACHE II than the approach that is normal GCS for sedated patients^{(24)}. Inclusion criteria of a study can also lead to calibration problems. Coronary care and post-cardiac surgical patients were included in the present study; however, the lack of goodnessof-fit for both models is shown persistently after excluding these populations. The number and type of missing physiological variable may affect the prediction of mortality⁽²⁵⁾. In the original APACHE II, physiologic variable were missing in $13%$ of the cases⁽⁴⁾. Engel et al (26) reported bilirubin, oxygenation, and GCS to be missing in 84%, 34%, and 34% of cases, respectively, in SAPS II database. In the presented records, missing physiological variables were found in only 6% for APACHE II requirements and 6.3% for SAPS II requirement (excluding bilirubin, which was missing in 76.5% of the presented data records).

The potential role of difference in case mix between the presented database and the development database may have had a negative impact of calibration assessment. In general, medical patients have a higher mortality risk than postoperative surgical patients, and in the present study population, medical patients constituted a larger proportion (55.6%) than in the original SAPS II database $(48\%)^{(6)}$. However, when the main diagnostic categories in the presented population are compared with those in the development APACHE II, some interesting differences appear. "Sepsis", diagnostic of "cardiovascular" without one of these principle diagnosis in non-operative patients and "gastrointestinal perforation/obstruction" in postoperative patients which are associated with a high mortality risk account for 16.7%, 3.2% and 3.4% of the presented database, compared with 3.6%, 1.1% and 2.1% in the original APACHE $II^{(4)}$, respectively. On the other hand, the postoperative category of "peripheral vascular surgery", which is associated with a low mortality rate only accounted for 0.8% of our admissions, compared with 9.8% of the development database⁽⁴⁾. These differences could contribute to overpredicted death for the APACHE II model in the presented patients.

Lead time bias is another factor that could affect the accuracy of hospital prediction. Tunnell et al (27) revealed that lead time bias increased the APACHE II and SAPS II scores by 14 and 23 points, respectively, leading to the APACHE II and SAPS II prediction of hospital mortality being increased as much as 42.7% and 33.4%, respectively. In the present study, however, the lead time was no different between survivors and non-survivors, and univariate analysis showed that lead time was not a significant prediction of hospital deaths (odds ratio $1.01, 95\%$ CI = 0.99-1.02 per day, $p = 0.16$). Thus, the authors believed that the influence of lead time bias on calibrations of both models is minimal in the present study.

The present study has some limitations. First, as a single center study there may be bias concerning case mix, quality of ICU care, and ICU policy. Secondly, the relatively small sample size was a relevant limiting factor in performed stratified analysis of calibration of both models. A multicenter study would have the benefit of fewer concerns of case mix and a better sample size.

In conclusion, the present study found two popular severity scoring systems in a Thai ICU to be less accurate than in the original studies. Both models showed poor calibration but excellent discrimination power, although overall the APACHE II performed better than the SAPS II model. A locally customized APACHE II, or a new version of a scoring system such as APACHE III or SAPS III, should also be evaluated in Thai ICUs.

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เปรียบเทียบระบบคะแนนความรุนแรง APACHE II และ SAPS II ในการทำนายอัตราตายใน โรงพยาบาลของผใหญู้ ชาวไทยในหออภ ่ บาลผ ิ ปู้วย่

บดนทร ิ ์ ขวญน ั มิตริ , อลนั กเตอร ี ์

วตถั ประสงค ุ :์ เพอประเม ่ื นความสามารถของระบบ ิ APACHE II และ SAPS II ในผปู้ วยหน ่ กในประเทศไทย ั **วัสดุและวิธีการ**: ศึกษาแบบติดตามไปข[้]างหน[้]าในช[่]วง 1 กรกฎาคม พ.ศ. 2547 ถึง 31 ตุลาคม พ.ศ. 2548 ในหออภิบาล *ผู้ป่วยโรงพยาบาลสงขลานครินทร์ซึ่งเป็นโรงพยาบาลมหาวิทยาลัยระดับตติยภูมิขนาด 800 เตียง*

ผลการศึกษา: ผู้ป[่]วยทั้งหมด 1,316 ราย เสียชีวิตในโรงพยาบาล 310 ราย (ร[้]อยละ 23.6) ระบบ APACHE II และ *SAPS II ทำนายอตราตายในโรงพยาบาลร ั อยละ ้ 30.5 + 28.2 และ 30.5 + 29.8 ตามลำดบั ระบบคะแนนทงสอง ้ั มีค่าการจำแนกที่ดีมาก โดยระบบคะแนน APACHE II สามารถจำแนกผู้ป่วยรอดและเสียชีวิตได้ดีกว่าระบบ SAPS II (0.911 และ 0.888, p < 0.001) อย่างไรก็ตามระบบทั้งสองไม่สามารถทำนายอัตราตายได้เที่ยงตรง Hosmer-Lemeshow goodness-of-fit H และ C statistics เทาก่ บั 66.59 และ 66.65 สำหรบระบบ ั APACHE II (p < 0.001) และ 54.01 และ 71.44 สำหรบระบบ ั SAPS II (p < 0.001)*

สรุป: ระบบ APACHE II จำแนกผู้ป่วยรอดและเสียชีวิตได้ดีกว่าระบบ SAPS II แต่ระบบทั้งสองทำนายอัตราตาย ไม่เที่ยงตรงโดยทำนายสูงกว่าอัตราตายจริง ควรปรับปรุงหรือหาระบบคะแนนความรุนแรงใหม่ที่เหมาะสมสำหรับ ผู้ป่วยหนักในประเทศไทย